

Source of the Data and Accuracy of the Estimates for the August 2022 Current Population Survey Microdata File on Veterans

SOURCE OF THE DATA

The data in this microdata file are from the August 2022 Current Population Survey (CPS). The U.S. Census Bureau conducts the CPS every month, although this file has only August data. The August survey uses two sets of questions, the basic CPS and a set of supplemental questions. The CPS, sponsored jointly by the Census Bureau and the U.S. Bureau of Labor Statistics, is the country's primary source of labor force statistics for the civilian noninstitutionalized population. The Department of Veterans Affairs and the Department of Labor jointly sponsor the supplemental questions for August.

Basic CPS. The monthly CPS collects primarily labor force data about the civilian noninstitutionalized population living in the United States. The institutionalized population, which is excluded from the universe, consists primarily of the population in correctional institutions and nursing homes (98 percent of the 4.0 million institutionalized people in the 2010 Census). Starting in August 2017, college and university dormitories were also excluded from the universe because most of the residents had usual residences elsewhere. Interviewers ask questions concerning labor force participation of each member 15 years old and older in sample households. Typically, the week containing the nineteenth of the month is the interview week. The week containing the twelfth is the reference week (i.e., the week about which the labor force questions are asked).

The CPS uses a multistage probability sample based on the results of the decennial census, with coverage in all 50 states and the District of Columbia. The sample is continually updated to account for new residential construction. When files from the most recent decennial census become available, the Census Bureau gradually introduces a new sample design for the CPS.

Every ten years, the CPS first-stage sample is redesigned¹ reflecting changes based on the most recent decennial census. In the first stage of the sampling process, primary sampling units (PSUs)² were selected for sample. In the 2010 sample design, the United States was divided into 1,987 PSUs. These PSUs were then grouped into 852 strata. Within each stratum, a single PSU was chosen for the sample, with its probability of selection proportional to its population as of the most recent decennial census. In the case of strata consisting of only one PSU, the PSU was chosen with certainty.

Approximately 68,500 sampled addresses were selected from the sampling frame in August. Based on eligibility criteria, eight percent of these sampled addresses were sent directly to computer-assisted telephone interviewing (CATI). The remaining sampled

¹ For detailed information on the 2010 sample redesign, please reference Bureau of Labor Statistics (2014).

² The PSUs correspond to substate areas (i.e., counties or groups of counties) that are geographically contiguous.

addresses were assigned to interviewers for computer-assisted personal interviewing (CAPI).³ Of all addresses in sample, about 59,000 were determined to be eligible for interview. Interviewers obtained interviews at about 43,500 of the housing units at these addresses. Noninterviews occur when the occupants are not found at home after repeated calls or are unavailable for some other reason.

August 2022 Supplement. In August 2022, in addition to the basic CPS questions, interviewers asked supplementary questions of veterans on year of discharge, disability, and job assistance.

Estimation Procedure. This survey's estimation procedure adjusts weighted sample results to agree with independently derived population controls of the civilian noninstitutionalized population of the United States, each state, and the District of Columbia. These population controls⁴ are prepared monthly as part of the Census Bureau's Population Estimates Program.

The population controls for the nation are distributed by demographic characteristics in two ways:

- Age, sex, and race (White alone, Black alone, and all other groups combined).
- Age, sex, and Hispanic origin.

The population controls for the states are distributed by:

- Race (Black alone and all other race groups combined).
- Age (0-15, 16-44, and 45 and over).
- Sex.

The independent estimates by age, sex, race, and Hispanic origin, and for states by selected age groups and broad race categories, are developed using the basic demographic accounting formula whereby the population from the 2020 Census data is updated using data on the components of population change (births, deaths, and net international migration) with net internal migration as an additional component in the state population controls.

The net international migration component of the population controls includes:

- Net international migration of the foreign born;
- Net migration between the United States and Puerto Rico;
- Net migration of natives to and from the United States; and

³ For further information on CATI and CAPI and the eligibility criteria, please reference U.S. Census Bureau (2019).

⁴ For additional information on population controls, including details on the demographic characteristics used and net international components, please refer to Chapters 1-3 and Appendix: History of the Current Population Survey of U.S. Census Bureau (2019).

- Net movement of the Armed Forces population to and from the United States.

Because the latest available information on these components lags behind the survey date, it is necessary to make short-term projections of these components to develop the estimate for the survey date.

ACCURACY OF THE ESTIMATES

A sample survey estimate has two types of error: sampling and nonsampling. The accuracy of an estimate depends on both types of error. The nature of the sampling error is known given the survey design; the full extent of the nonsampling error is unknown.

Sampling Error. Since the CPS estimates come from a sample, they may differ from figures from an enumeration of the entire population using the same questionnaires, instructions, and enumerators. For a given estimator, the difference between an estimate based on a sample and the estimate that would result if the sample were to include the entire population is known as sampling error. Standard errors, as calculated by methods described in “Standard Errors and Their Use,” are primarily measures of the magnitude of sampling error. However, the estimation of standard errors may include some nonsampling error.

Nonsampling Error. For a given estimator, the difference between the estimate that would result if the sample were to include the entire population and the true population value being estimated is known as nonsampling error. There are several sources of nonsampling error that may occur during the development or execution of the survey. It can occur because of circumstances created by the interviewer, the respondent, the survey instrument, or the way the data are collected and processed. Some nonsampling errors, and examples of each, include:

- Measurement error: The interviewer records the wrong answer, the respondent provides incorrect information, the respondent estimates the requested information, or an unclear survey question is misunderstood by the respondent.
- Coverage error: Some individuals who should have been included in the survey frame were missed.
- Nonresponse error: Responses are not collected from all those in the sample or the respondent is unwilling to provide information.
- Imputation error: Values are estimated imprecisely for missing data.
- Processing error: Forms may be lost, data may be incorrectly keyed, coded, or recoded, etc.

To minimize these errors, the Census Bureau applies quality control procedures during all stages of the production process including the design of the survey, the wording of questions, the review of the work of interviewers and coders, and the statistical review of reports.

Two types of nonsampling error that can be examined to a limited extent are nonresponse and undercoverage.

Nonresponse. The effect of nonresponse cannot be measured directly, but one indication of its potential effect is the nonresponse rate. For the August 2022 basic CPS, the household-level unweighted nonresponse rate was 26.3 percent. The person-level unweighted nonresponse rate for the Veterans supplement was an additional 16.2 percent. Since the basic CPS nonresponse rate is a household-level rate and the Veterans supplement nonresponse rate is a person-level rate, we cannot combine these rates to derive an overall nonresponse rate. Nonresponding households may have more or fewer persons than interviewed ones, so combining these rates may lead to an under- or overestimate of the true overall nonresponse rate for persons for the Veterans supplement.

Responses are made up of complete interviews and sufficient partial interviews. A sufficient partial interview is an incomplete interview in which the household or person answered enough of the questionnaire for the supplement sponsor to consider the interview complete. The remaining supplement questions may have been edited or imputed to fill in missing values. Insufficient partial interviews are considered to be nonrespondents. Refer to the supplement overview attachment in the technical documentation for the specific questions deemed critical by the sponsor as necessary to answer in order to be considered a sufficient partial interview.

As a result of sufficient partial interviews being considered responses, individual items/questions have their own response and refusal rates. As part of the nonsampling error analysis, the item response rates, item refusal rates, and edits are reviewed. For the Veterans supplement, the unweighted item refusal rates range from 0.0 percent to 4.4 percent. The unweighted item allocation rates range from 4.0 percent to 5.2 percent. The unweighted item nonresponse rates range from 0.0 percent to 11.6 percent.

Undercoverage. The concept of coverage with a survey sampling process is defined as the extent to which the total population that could be selected for sample “covers” the survey’s target population. Missed housing units and missed people within sample households create undercoverage in the CPS. Overall CPS undercoverage for August 2022 is estimated to be about ten percent. CPS coverage varies with age, sex, and race. Generally, coverage is higher for females than for males and higher for non-Blacks than for Blacks. This differential coverage is a general problem for most household-based surveys.

The CPS weighting procedure mitigates bias from undercoverage, but biases may still be present when people who are missed by the survey differ from those interviewed in ways other than age, race, sex, Hispanic origin, and state of residence. How this weighting procedure affects other variables in the survey is not precisely known. All of these considerations affect comparisons across different surveys or data sources.

A common measure of survey coverage is the coverage ratio, calculated as the estimated population before poststratification divided by the independent population control. Table 1 shows August 2022 CPS coverage ratios by age and sex for certain race and Hispanic groups. The CPS coverage ratios can exhibit some variability from month to month.

Table 1. Current Population Survey Coverage Ratios: August 2022

| Age group | Total | | | White alone | | Black alone | | Residual race ^A | | Hispanic ^B | |
|-----------|------------|------|--------|-------------|--------|-------------|--------|----------------------------|--------|-----------------------|--------|
| | All people | Male | Female | Male | Female | Male | Female | Male | Female | Male | Female |
| 0-15 | 0.85 | 0.86 | 0.83 | 0.91 | 0.87 | 0.72 | 0.68 | 0.79 | 0.78 | 0.83 | 0.81 |
| 16-19 | 0.82 | 0.84 | 0.80 | 0.86 | 0.84 | 0.77 | 0.68 | 0.78 | 0.73 | 0.89 | 0.82 |
| 20-24 | 0.76 | 0.76 | 0.77 | 0.79 | 0.79 | 0.58 | 0.70 | 0.77 | 0.72 | 0.80 | 0.77 |
| 25-34 | 0.82 | 0.81 | 0.82 | 0.87 | 0.85 | 0.57 | 0.65 | 0.77 | 0.86 | 0.80 | 0.83 |
| 35-44 | 0.88 | 0.86 | 0.91 | 0.88 | 0.93 | 0.68 | 0.77 | 0.89 | 0.90 | 0.79 | 0.93 |
| 45-54 | 0.89 | 0.87 | 0.90 | 0.89 | 0.93 | 0.76 | 0.80 | 0.90 | 0.80 | 0.78 | 0.87 |
| 55-64 | 0.92 | 0.89 | 0.95 | 0.91 | 0.97 | 0.80 | 0.85 | 0.88 | 0.93 | 0.75 | 0.87 |
| 65+ | 1.06 | 1.04 | 1.07 | 1.06 | 1.08 | 1.00 | 1.02 | 0.90 | 0.94 | 0.88 | 0.90 |
| 15+ | 0.90 | 0.89 | 0.92 | 0.91 | 0.95 | 0.73 | 0.80 | 0.85 | 0.86 | 0.81 | 0.86 |
| 0+ | 0.89 | 0.88 | 0.90 | 0.91 | 0.93 | 0.73 | 0.77 | 0.83 | 0.84 | 0.81 | 0.85 |

Source: U.S. Census Bureau, Current Population Survey, August 2022.

^A The Residual race group includes cases indicating a single race other than White or Black, and cases indicating two or more races.

^B Hispanics may be any race.

Note: For a more detailed discussion on the use of parameters for race and ethnicity, please refer to the “Generalized Variance Parameters” section.

Comparability of Data. Data obtained from the CPS and other sources are not entirely comparable. This is due to differences in interviewer training and experience and in differing survey processes.⁵ These differences are examples of nonsampling variability not reflected in the standard errors. Therefore, caution should be used when comparing results from different sources.

Data users should be careful when comparing the data from this microdata file, which reflects 2020 Census-based controls⁶, with microdata files which reflect 2010 Census-based controls. Ideally, the same population controls should be used when comparing any estimates. In reality, the use of the same population controls is not practical when

⁵ Survey processes include, but are not limited to, question wording, universe, sampling frame, interview modes, and weighting.

⁶ In recent decades, the decennial census has usually provided all the data necessary to produce the population base used in the population controls. However, changes in disclosure avoidance practices and delays in the 2020 Census necessitated changes to the data sources that produce the base population for the Vintage 2021 population estimates. The updated population controls use a Blended Base that draws on the 2020 Census, 2020 Demographic Analysis Estimates, and Vintage 2020 Postcensal Population Estimates. More information on this methodology can be found at <<https://www2.census.gov/programs-surveys/popest/technical-documentation/methodology/2020-2021/methods-statement-v2021.pdf>>.

comparing trend data over a period of 10 to 20 years. Thus, when it is necessary to combine or compare data based on different controls or different designs, data users should be aware that changes in weighting controls or weighting procedures can create small differences between estimates. The discussion following includes information on comparing estimates derived from different populations or different sample designs.

Microdata files from previous years reflect the latest available census-based controls. Although the most recent change in population controls had relatively little impact on summary measures such as averages, medians, and percentage distributions, it did have a significant impact on levels. For example, use of 2020 Census-based controls results in about a 0.7 percent increase from the 2010 Census-based controls in the civilian noninstitutionalized population. Thus, estimates of levels for data collected in 2012 and later years will differ from those for earlier years by more than what could be attributed to actual changes in the population. These differences could be disproportionately greater for certain population subgroups than for the total population.

Users should also exercise caution because of changes caused by the phase-in of the 2010 Census files (refer to “Basic CPS”).⁷ During this time period, CPS data were collected from sample designs based on different censuses. Two features of the new CPS design have the potential of affecting estimates: (1) the temporary disruption of the rotation pattern from August 2014 through June 2015 for a comparatively small portion of the sample and (2) the change in sample areas. Most of the known effect on estimates during and after the sample redesign will be the result of changing from 2000 to 2010 geographic definitions. Research has shown that the national-level estimates of the metropolitan and nonmetropolitan populations should not change appreciably because of the new sample design. However, users should still exercise caution when comparing metropolitan and nonmetropolitan estimates across years with a design change, especially at the state level.

Caution should also be used when comparing Hispanic estimates over time. No independent population control totals for people of Hispanic origin were used before 1985.

A Nonsampling Error Warning. Since the full extent of the nonsampling error is unknown, one should be particularly careful when interpreting results based on small differences between estimates. The Census Bureau recommends that data users incorporate information about nonsampling errors into their analyses, as nonsampling error could impact the conclusions drawn from the results. Caution should also be used when interpreting results based on a relatively small number of cases. Summary measures (such as medians and percentage distributions) probably do not reveal useful information when computed on a subpopulation smaller than 75,000.

For additional information on nonsampling error, including the possible impact on CPS data, when known, refer to U.S. Census Bureau (2019) and Brooks & Bailar (1978).

⁷ The phase-in process using the 2010 Census files began April 2014.

Standard Errors and Their Use. A sample estimate and its standard error enable one to construct a confidence interval. A confidence interval is a range about a given estimate that has a specified probability of containing the average result of all possible samples. For example, if all possible samples were surveyed under essentially the same general conditions and using the same sample design, and if an estimate and its standard error were calculated from each sample, then approximately 90 percent of the intervals from 1.645 standard errors below the estimate to 1.645 standard errors above the estimate would include the average result of all possible samples.

A particular confidence interval may or may not contain the average estimate derived from all possible samples, but one can say with the specified confidence that the interval includes the average estimate calculated from all possible samples.

Standard errors may also be used to perform hypothesis testing, a procedure for distinguishing between population parameters using sample estimates. The most common type of hypothesis is that the population parameters are different. An example of this would be comparing the percentage of men who were part-time workers to the percentage of women who were part-time workers.

Tests may be performed at various levels of significance. A significance level is the probability of concluding that the characteristics are different when, in fact, they are the same. For example, to conclude that two characteristics are different at the 0.10 level of significance, the absolute value of the estimated difference between characteristics must be greater than or equal to 1.645 times the standard error of the difference.

The Census Bureau uses 90-percent confidence intervals and 0.10 levels of significance to determine statistical validity. Consult standard statistical textbooks for alternative criteria.

Estimating Standard Errors. The Census Bureau uses replication methods to estimate the standard errors of CPS estimates. These methods primarily measure the magnitude of sampling error. However, they do measure some effects of nonsampling error as well. They do not measure systematic biases in the data associated with nonsampling error. Bias is the average over all possible samples of the differences between the sample estimates and the true value.

Generalized Variance Parameters. While it is possible to estimate the standard error based on the survey data for each estimate in a report, there are a number of reasons why this is not done. A presentation of the individual standard errors would be of limited use, since one could not possibly predict all of the combinations of results that may be of interest to data users. Additionally, data users have access to CPS microdata files, and it is impossible to compute in advance the standard error for every estimate one might obtain from those data sets. Moreover, variance estimates are based on sample data and have variances of their own. Therefore, some methods of stabilizing these estimates of variance, for example, by generalizing or averaging over time, may be used to improve their reliability.

Experience has shown that certain groups of estimates have similar relationships between their variances and expected values. Modeling or generalizing may provide more stable variance estimates by taking advantage of these similarities. The generalized variance function (GVF) is a simple model that expresses the variance as a function of the expected value of the survey estimate. The parameters of the GVF are estimated using direct replicate variances. These GVF parameters provide a relatively easy method to obtain approximate standard errors for numerous characteristics.

In this source and accuracy statement:

- Tables 3 through 5 provide illustrations for calculating standard errors;
- Table 6 provides the GVF parameters for labor force estimates; and
- Table 7 provides GVF parameters for characteristics from the August 2022 supplement.

The basic CPS questionnaire records the race and ethnicity of each respondent. With respect to race, a respondent can be White, Black, Asian, American Indian and Alaskan Native (AIAN), Native Hawaiian and Other Pacific Islander (NHOPI), or combinations of two or more of the preceding. A respondent's ethnicity can be Hispanic or non-Hispanic, regardless of race.

The GVF parameters to use in computing standard errors are dependent upon the race/ethnicity group of interest. Table 2 summarizes the relationship between the race/ethnicity group of interest and the GVF parameters to use in standard error calculations.

Table 2. Estimation Groups of Interest and Generalized Variance Parameters

| Race/ethnicity group of interest | Generalized variance parameters to use in standard error calculations |
|---|---|
| Total population | Total or White |
| White alone, White alone or in combination (AOIC), or White non-Hispanic population | Total or White |
| Black alone, Black AOIC, or Black non-Hispanic population | Black |
| Asian alone, Asian AOIC, or Asian non-Hispanic population | Asian, American Indian and Alaska Native (AIAN), Native Hawaiian and Other Pacific Islander (NHOPI) |
| AIAN alone, AIAN AOIC, or AIAN non-Hispanic population | Asian, AIAN, NHOPI |
| NHOPI alone, NHOPI AOIC, or NHOPI non-Hispanic population | Asian, AIAN, NHOPI |
| Populations from other race groups | Asian, AIAN, NHOPI |
| Hispanic ^A population | Hispanic ^A |
| Two or more races ^B – employment/unemployment and educational attainment characteristics | Black |
| Two or more races ^B – all other characteristics | Asian, AIAN, NHOPI |

Source: U.S. Census Bureau, Current Population Survey, internal data files.

^A Hispanics may be any race.

^B Two or more races refers to the group of cases self-classified as having two or more races.

When calculating standard errors for an estimate of interest from cross-tabulations involving different characteristics, use the set of GVF parameters for the characteristic that will give the largest standard error. If the estimate of interest is strictly from basic CPS data, the GVF parameters will come from the CPS GVF table (Table 6). If the estimate is using Veterans supplement data, the GVF parameters will come from the Veterans supplement GVF table (Table 7).

Standard Errors of Estimated Numbers. The approximate standard error, s_x , of an estimated number from this microdata file can be obtained by using the formula:

$$s_x = \sqrt{ax^2 + bx} \tag{1}$$

Here x is the size of the estimate, and a and b are the parameters in Table 6 or 7 associated with the particular type of characteristic.

Illustration 1

Suppose there were 2,933,000 unemployed nonveterans, aged 18 to 34, in the civilian labor force. Table 3 shows how to use the appropriate parameters from Table 6 and Formula (1) to estimate the standard error and confidence interval.

Table 3. Illustration of Standard Errors of Estimated Numbers

| | |
|--|------------------------|
| Number of unemployed nonveterans in the civilian labor force (x) | 2,933,000 |
| a-parameter (a) | -0.000017 |
| b-parameter (b) | 3,244 |
| Standard error | 97,000 |
| 90-percent confidence interval | 2,774,000 to 3,093,000 |

Source: U.S. Census Bureau, Current Population Survey, August 2022.

The standard error is calculated as

$$s_x = \sqrt{-0.000017 \times 2,933,000^2 + 3,244 \times 2,933,000},$$

which, rounded to the nearest thousand, is 97,000. The 90-percent confidence interval is calculated as $2,933,000 \pm 1.645 \times 97,000$.

A conclusion that the average estimate derived from all possible samples lies within a range computed in this way would be correct for roughly 90 percent of all possible samples.

Standard Errors of Estimated Percentages. The reliability of an estimated percentage, computed using sample data for both numerator and denominator, depends on both the size of the percentage and its base. Estimated percentages are relatively more reliable than the corresponding estimates of the numerators of the percentages, particularly if the percentages are 50 percent or more. When the numerator and denominator of the percentage are in different categories, use the parameter from Table 6 or 7 as indicated by the numerator.

The approximate standard error, $s_{y,p}$, of an estimated percentage can be obtained by using the formula:

$$s_{y,p} = \sqrt{\frac{b}{y} p(100 - p)} \quad (2)$$

Here y is the total number of people, families, households, or unrelated individuals in the base or denominator of the percentage, p is the percentage $100 \cdot x/y$ ($0 \leq p \leq 100$), and b is the parameter in Table 6 or 7 associated with the characteristic in the numerator of the percentage.

Illustration 2

Suppose there were 1,424,000 Gulf War veterans aged 18 to 34 in the civilian labor force, and 2.7 percent were unemployed. Table 4 shows how to use the appropriate parameters from Table 7 and Formula (2) to estimate the standard error and confidence interval.

Table 4. Illustration of Standard Errors of Estimated Percentages

| | |
|---|------------|
| Percentage of Gulf War veterans, aged 18-34, unemployed (p) | 2.7 |
| Base (y) | 1,424,000 |
| b-parameter (b) | 5,011 |
| Standard error | 0.96 |
| 90-percent confidence interval | 1.1 to 4.3 |

Source: U.S. Census Bureau, Current Population Survey, Veterans, August 2022.

The standard error is calculated as

$$s_{y,p} = \sqrt{\frac{5,011}{1,424,000} \times 2.7 \times (100.0 - 2.7)} = 0.96$$

and the 90-percent confidence interval for the estimated percentage of unemployed Gulf War veterans aged 18 to 34 in the civilian labor force is from 1.1 to 4.3 percent (i.e., $2.7 \pm 1.645 \times 0.96$).

Standard Errors of Estimated Differences. The standard error of the difference between two sample estimates is approximately equal to

$$s_{|x_1-x_2|} = \sqrt{(s_{x_1})^2 + (s_{x_2})^2} \quad (3)$$

Where s_{x_1} and s_{x_2} are the standard errors of the estimates, x_1 and x_2 . The estimates can be numbers, percentages, ratios, etc. This will result in accurate estimates of the standard error of the same characteristic in two different areas or for the difference between separate and uncorrelated characteristics in the same area. However, if there is a high positive (negative) correlation between the two characteristics, the formula will overestimate (underestimate) the true standard error.

Illustration 3

Suppose that of the 1,424,000 Gulf War veterans in the civilian labor force between 18 and 34 years of age, 2.7 percent were unemployed, and of the 54,510,000 nonveterans in the civilian labor force between 18 and 34 years of age, 5.4 percent were unemployed. Table 5 shows how to use the appropriate parameters from Tables 6 and 7 and Formulas (2) and (3) to estimate the standard error and confidence interval.

Table 5. Illustration of Standard Errors of Estimated Differences

| | Gulf War Veterans (x_1) | Nonveterans (x_2) | Difference |
|---|-----------------------------|-----------------------|------------|
| Percentage unemployed aged 18 to 34 (p) | 2.7 | 5.4 | 2.7 |
| Base (y) | 1,424,000 | 54,510,000 | - |
| b-parameter (b) | 5,011 | 3,244 | - |
| Standard error | 0.96 | 0.17 | 0.97 |
| 90-percent confidence interval | 1.1 to 4.3 | 5.1 to 5.7 | 1.1 to 4.3 |

Source: U.S. Census Bureau, Current Population Survey, Veterans, August 2022.

The standard error of the difference is calculated as

$$s_{|x_1-x_2|} = \sqrt{0.96^2 + 0.17^2} = 0.97$$

and the 90-percent confidence interval around the difference is calculated as $2.7 \pm 1.645 \times 0.97$. Since this interval does not include zero, we can conclude with 90-percent confidence that the percentage of unemployed Gulf War veterans in the civilian labor force between 18 and 34 years of age is lower than the percentage of unemployed nonveterans in the civilian labor force between 18 and 34 years of age.

Standard Errors of Quarterly or Yearly Averages. For information on calculating standard errors for labor force data from the CPS which involve quarterly or yearly averages, please reference Bureau of Labor Statistics (2006).

Technical Assistance. If you require assistance or additional information, please contact the Demographic Statistical Methods Division via e-mail at dsmd.source.and.accuracy@census.gov.

Table 6. Parameters for Computation of Standard Errors for Labor Force Characteristics: August 2022

| Characteristic | a | b |
|--|-----------|----------|
| Total or White | | |
| <i>Civilian labor force, employed</i> | -0.000013 | 2,481 |
| <i>Unemployed</i> | -0.000017 | 3,244 |
| <i>Not in labor force</i> | -0.000013 | 2,432 |
| <i>Civilian labor force, employed, not in labor force, and unemployed</i> | | |
| Men | -0.000031 | 2,947 |
| Women | -0.000028 | 2,788 |
| Both sexes, 16 to 19 years | -0.000261 | 3,244 |
| Black | | |
| <i>Civilian labor force, employed, not in labor force, and unemployed</i> | | |
| Total | -0.000117 | 3,601 |
| Men | -0.000249 | 3,465 |
| Women | -0.000191 | 3,191 |
| Both sexes, 16 to 19 years | -0.001425 | 3,601 |
| Asian, American Indian and Alaska Native (AIAN), Native Hawaiian and Other Pacific Islander (NHOPI) | | |
| <i>Civilian labor force, employed, not in labor force, and unemployed</i> | | |
| Total | -0.000245 | 3,311 |
| Men | -0.000537 | 3,397 |
| Women | -0.000399 | 2,874 |
| Both sexes, 16 to 19 years | -0.004078 | 3,311 |
| Hispanic, may be of any race | | |
| <i>Civilian labor force, employed, not in labor force, and unemployed</i> | | |
| Total | -0.000087 | 3,316 |
| Men | -0.000172 | 3,276 |
| Women | -0.000158 | 3,001 |
| Both sexes, 16 to 19 years | -0.000909 | 3,316 |

Source: U.S. Census Bureau, Internal Current Population Survey data files for the 2010 Design.

Notes: These parameters are to be applied to basic CPS monthly labor force estimates. The Total or White, Black, and Asian, AIAN, NHOPI parameters are to be used for both alone and in combination race group estimates. For nonmetropolitan characteristics, multiply the a- and b-parameters by 1.5. If the characteristic of interest is total state population, not subtotaled by race or ethnicity, the a- and b-parameters are zero. For foreign-born and noncitizen characteristics for Total and White, the a- and b-parameters should be multiplied by 1.3. No adjustment is necessary for foreign-born and noncitizen characteristics for Black, Hispanic, and Asian, AIAN, NHOPI parameters. For the groups self-classified as having two or more races, use the Asian, AIAN, NHOPI parameters for all employment characteristics.

**Table 7. Parameters for Computation of Standard Errors for Veterans Characteristics:
August 2022**

| Characteristic | a | b |
|--|-----------|-------|
| Total Employed and Nonagriculture Employed in Labor Force, Occupations, and Disability Status of Employed | | |
| Total or Men | | |
| All Veterans | -0.000266 | 4,885 |
| War Veterans | -0.000339 | 4,885 |
| Gulf War Era Veterans | -0.000599 | 4,885 |
| Other Service Veterans | -0.001228 | 4,885 |
| Women | | |
| All Veterans | -0.002408 | 4,885 |
| War Veterans | -0.003015 | 4,885 |
| Gulf War Era Veterans | -0.003558 | 4,885 |
| Other Service Veterans | -0.011944 | 4,885 |
| Unemployed, Duration of Unemployment | | |
| Total or Men | | |
| All Veterans | -0.000272 | 5,011 |
| War Veterans | -0.000347 | 5,011 |
| Gulf War Era Veterans | -0.000614 | 5,011 |
| Other Service Veterans | -0.001260 | 5,011 |
| Women | | |
| All Veterans | -0.002470 | 5,011 |
| War Veterans | -0.003093 | 5,011 |
| Gulf War Era Veterans | -0.003650 | 5,011 |
| Other Service Veterans | -0.012252 | 5,011 |

Source: U.S. Census Bureau, Current Population Survey, Internal data from the Veterans Supplement, August 2022.

Notes: These parameters are to be applied to the Veterans Supplement data. For foreign-born characteristics, the a- and b-parameters should be multiplied by 1.3.

REFERENCES

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